

Lecture Notes on Hybrid Uncertainties

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Source material:

- Yakov Ben-Haim, 2006, *Info-Gap Decision Theory: Decisions Under Severe Uncertainty*, 2nd edition, Academic Press, chapter 10.

- Yakov Ben-Haim, 1996, *Robust Reliability in the Mechanical Sciences*, Springer, chap. 8.

A Note to the Student: These lecture notes are not a substitute for the thorough study of books. These notes are no more than an aid in following the lectures.

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- ¶ Sometimes one has both **probabilistic** and **info-gap** information about the uncertainties.
- ¶ Neither is sufficient to fully characterize the uncertainty.
- ¶ We will consider three situations:
 - Info-gap uncertainty and the Poisson process.
 - Uncertain probability distributions embedded in an info-gap model.
 - Probabilistic info-gap horizon of uncertainty.

1 Info-Gap Uncertainty in a Poisson Process

1.1 Poisson and Info-Gap Uncertainties

¶ Many complex events such as earthquakes, currency crashes, or other extreme disturbances have **two distinct time constants**:

1. The events recur infrequently over time.

That is, on the **long time scale**, θ , they can be thought of as distinct points.

2. The temporal variation during an event is both important and unknown.

That is, on the **short time scale**, t , they are complex and unknown.

¶ A common and often reliable statistical datum on the long time scale is: Average rate of recurrence of a rare event over a long duration θ .

¶ The **poisson process** is a good probabilistic model for long durations if:

1. The occurrence of distinct events is statistically independent.
2. The average number of events per unit of time is constant.

¶ With these two assumptions, the probability of exactly n events in a duration θ is given by the Poisson distribution:

$$P_n(\theta) = \frac{(\lambda\theta)^n e^{-\lambda\theta}}{n!}, \quad n = 0, 1, 2, \dots \quad (1)$$

¶ This is valid for representing distributions in space as well as in time.

¶ The mean number of events in duration θ is:

$$E[n(\theta)] = \lambda\theta \quad (2)$$

¶ Thus λ = mean number of events per unit time.

¶ An info-gap model is a good representation of the uncertain variation of the temporal waveform during an event.

1.2 Shock Loading of a Dynamical System

¶ Dynamical system:

- t = short time scale.
- $x_u(t)$ = state vector.
- $u(t)$ = Severe transient load vector.

¶ Damage due to loads:

- Severe loads recur infrequently, causing damage.
- Damage depends on the short-time-scale dynamics.
- Damage accumulates from each event, until the system fails.

¶ System model:

$$\frac{dx}{dt} = Ax(t) + Bu(t), \quad x(0) = 0 \quad (3)$$

A and B are known constant matrices.

¶ Cumulative energy-bound load-uncertainty model:

$$\mathcal{U}(h, \tilde{u}) = \left\{ u(t) : \int_0^\infty [u(t) - \tilde{u}(t)]^T W [u(t) - \tilde{u}(t)] dt \leq h^2 \right\}, \quad h \geq 0 \quad (4)$$

W is a known, real, symmetric, positive definite matrix.

¶ Small increment of damage resulting from one event:

$$\delta_u = \gamma \left[\psi^T x_u(t) \right]^\mu \quad (5)$$

γ and μ are known, positive constants.

ψ is a known projection vector.

¶ Poisson probability, $P_n(\theta)$, of n transient events in a long interval of duration θ , eq.(1).
Single known parameter, λ .

¶ Failure occurs if the **cumulative damage** exceeds Δ_c .

1.3 Robustness Function: I

- ¶ Failure occurs in n events if the cumulative damage exceeds the critical value Δ_c .
- ¶ The robustness to $n > 0$ events, \hat{h}_n , is the greatest value of the uncertainty parameter h such that failure cannot occur in n events:

$$\hat{h}_n = \max \left\{ h : n \max_{u \in \mathcal{U}(h, \tilde{u})} \delta_u(t) \leq \Delta_c \right\} \quad (6)$$

We note that \hat{h}_n is meaningful for $n > 0$. Failure cannot occur if damage-inducing events do not occur.

1.4 Maximal Increment of Damage

¶ In order to evaluate the robustness function we must find the maximum increment of damage in a single event, up to uncertainty h .

¶ This requires the maximum projected response.

¶ The response to input $u(t)$ is:

$$x_u(t) = \int_0^t e^{A(t-\tau)} B u(\tau) d\tau \quad (7)$$

¶ The deviation of the projected response is:

$$\psi^T [x_u(t) - x_{\tilde{u}}(t)] = \int_0^t \psi^T e^{A(t-\tau)} B [u(\tau) - \tilde{u}(\tau)] d\tau \quad (8)$$

$$= \int_0^t \psi^T e^{A(t-\tau)} B W^{-1/2} W^{1/2} [u(\tau) - \tilde{u}(\tau)] d\tau \quad (9)$$

$$= \int_0^t \zeta^T(t-\tau) W^{1/2} [u(\tau) - \tilde{u}(\tau)] d\tau \quad (10)$$

where we have defined the vector:

$$\zeta^T(t) = \psi^T e^{At} B W^{-1/2} \quad (11)$$

¶ The maximum projected response up to uncertainty h is:

$$\max_{u \in \mathcal{U}(h, \tilde{u})} \psi^T [x_u(t) - x_{\tilde{u}}(t)] = h \underbrace{\sqrt{\int_0^t \zeta^T(\tau) \zeta(\tau) d\tau}}_{Z(t)} \quad (12)$$

which defines the known function $Z(t)$.

(Hint: use the Cauchy inequality, and then the Schwarz inequality.)

¶ Now, combining eqs.(5) and (12), the maximum increment of damage in a single transient event, up to uncertainty h , is:

$$\max_{u \in \mathcal{U}(h, \tilde{u})} \delta_u(t) = \gamma \left[\psi^T x_{\tilde{u}}(t) + h Z(t) \right]^\mu \quad (13)$$

1.5 Robustness Function: II

- ¶ Failure occurs in n events if the cumulative damage exceeds the critical value Δ_c .
- ¶ As explained in section 1.3, the robustness to $n > 0$ events, \hat{h}_n , is the greatest value of the uncertainty parameter h such that failure cannot occur in n events:

$$\hat{h}_n = \max \left\{ h : n \max_{u \in \mathcal{U}(h, \tilde{u})} \delta_u(t) \leq \Delta_c \right\} \quad (14)$$

We note that \hat{h}_n is meaningful for $n > 0$. Failure cannot occur if damage-inducing events do not occur.

- ¶ Equate max cumulative damage to Δ_c :

$$n \max_{u \in \mathcal{U}(h, \tilde{u})} \delta_u(t) = \Delta_c \quad (15)$$

Now solve for h to find the robustness to n transients:

$$\hat{h}_n = \frac{(\Delta_c/n\gamma)^{1/\mu} - \psi^T x_{\tilde{u}}(t)}{Z(t)}, \quad n = 1, 2, \dots \quad (16)$$

or $\hat{h}_n = 0$ if this is negative.

¶ n is a Poisson random variable.

Therefore \hat{h}_n is also a Poisson random variable.

¶ Randomization: concise combination of
info-gap and probabilistic information.

$$\hat{h}(\theta) = \frac{1}{1 - P_0(\theta)} \sum_{n=1}^{\infty} \hat{h}_n P_n(\theta) \quad (17)$$

We are usually interested in long durations θ for which:

$$P_0(\theta) = e^{-\lambda\theta} \ll 1 \quad (18)$$

¶ $\hat{h}(\theta)$ is a decision function, since “bigger is better”.

¶ Let q be the vector of decision variables. We will write $\hat{h}(q, \Delta_c)$.

¶ The optimal decision vector $\hat{q}_c(\Delta_c)$:

$$\hat{q}_c(\Delta_c) = \arg \max_{q \in \mathcal{Q}} \hat{h}(q, \Delta_c) \quad (19)$$

\mathcal{Q} = set of available decisions.

¶ Both robustness functions:

$\hat{h}(q, \Delta_c)$ and $\hat{h}(\hat{q}_c(\Delta_c), \Delta_c)$,

display the usual trade-off of immunity versus reward.

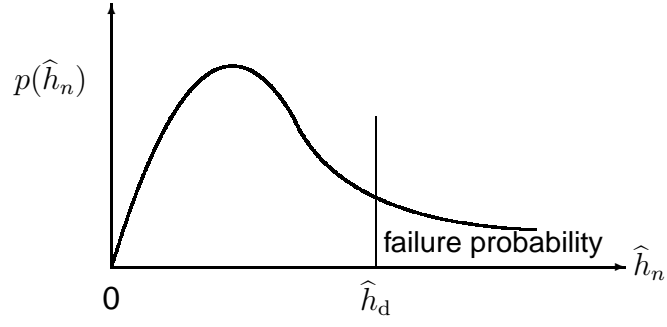


Figure 1: Illustration of failure probability for eq.(20).

- ¶ Different approach: Optimize probability distribution of \hat{h}_n .
- Let \hat{h}_d be a desired or demanded value of robustness.
 - Choose q to maximize the probability of those $\hat{h}_n(q)$'s which exceed the demanded value \hat{h}_d :

$$\hat{q}(\hat{h}_d) = \arg \max_{q \in \mathcal{Q}} \sum_{\hat{h}_n(q) \geq \hat{h}_d} P_n(\theta) \quad (20)$$

Let us examine the condition:

$$\hat{h}_n(q) \geq \hat{h}_d \quad (21)$$

From eq.(16) this becomes:

$$\left(\frac{\Delta_c}{n\gamma} \right)^{1/\mu} \geq \psi^T x_u^*(t) + \hat{h}_d Z(t) \quad (22)$$

Solving for n :

$$n \leq \frac{\Delta_c}{\gamma [\psi^T x_u^*(t) + \hat{h}_d Z(t)]^\mu} \quad (23)$$

We maximize the probability that condition (21) holds if we choose q to minimize $\psi^T x_u^*(t) + \hat{h}_d Z(t)$.

2 Embedded Probability Densities

¶ We consider the following situation:

- u is uncertain.
- The uncertainty in u is represented by a pdf $p(u)$.
- $p(u)$ is uncertain.
- The uncertainty in $p(u)$ is represented by an info-gap model.

2.1 Formulation: Dynamical System

¶ Variables:

- u = uncertain input to a system.
- x_u = response to input u .
- $p(u)$ = pdf of u ; imperfectly known.
- $\tilde{p}(u)$ = nominal pdf of u ; known.
- $\mathcal{U}(h, \tilde{p})$, $h \geq 0$: info-gap model for uncertainty of p .

¶ Failure occurs if:

$$f(x_u) > x_c \quad (24)$$

¶ For any pdf $p(u)$, the probability of failure is:

$$P_f(p) = \text{Prob}(f(x_u) > x_c | p) \quad (25)$$

$$= \int_{f(x_u) > x_c} p(u) du \quad (26)$$

¶ We want:

$$P_f(p) \leq P_c \quad (27)$$

¶ We cannot reliably calculate $P_f(p)$ because p is uncertain.

¶ We **can** calculate the robustness, to uncertainty in $p(u)$, of the failure probability:

$$\hat{h}(P_c) = \max \left\{ h : \max_{p \in \mathcal{U}(h, \tilde{p})} P_f(p) \leq P_c \right\} \quad (28)$$

This is an ordinary robustness function for uncertainty in p .

If $\hat{h}(P_c)$ is large then we have confidence, despite the info-gaps in the pdf, that the failure probability will not exceed P_c .

2.2 Example: 1-D Dynamic System

¶ 1-D system:

$$\frac{dx}{dt} = Ax(t) + Bu(t), \quad x(0) = 0 \quad (29)$$

A and B are known constant scalars.

¶ Variables:

- u = input.
- = constant random variable in $[0, T]$. Zero elsewhere.
- $p(u)$ = pdf of u .
- $\tilde{p}(u)$ = best-estimate of the probability density of u .
- = $\mathcal{N}(0, \sigma^2)$.

¶ Uncertainty in $p(u)$:

- Evidence for \tilde{p} is quite good up to about k standard deviations.
- Beyond $k\sigma$ the fractional deviation of p from \tilde{p} varies.
- An info-gap model for uncertainty in p is:

$$\mathcal{U}(h, \tilde{p}) = \left\{ p(u) : \begin{array}{l} p(u) \geq 0, \int p(u) du = 1, \\ |p(u) - \tilde{p}(u)| \leq h\tilde{p}(u) \text{ if } |u| \geq k\sigma \\ p(u) = c\tilde{p}(u) \text{ if } |u| < k\sigma \end{array} \right\}, \quad h \geq 0 \quad (30)$$

c is a normalization constant for each density $p(u)$.

¶ System response at end of nominal load:

$$x_u(T) = \frac{uB(e^{AT} - 1)}{A} \quad (31)$$

¶ Failure criterion:

$$|x_u(T)| > x_c \quad (32)$$

¶ Probability of failure, given density $p(u)$, is:

$$P_f(p) = \text{Prob}(|x_u(T)| > x_c | p) \quad (33)$$

$$= \text{Prob}(|u| > \eta x_c) \quad (34)$$

where we have defined:

$$\eta = \frac{A}{B(e^{AT} - 1)} \quad (35)$$

¶ As before, we desire:

$$P_f(p) \leq P_c \quad (36)$$

¶ Simplifying assumption:

$$\eta x_c \geq k\sigma \quad (37)$$

¶ To evaluate the robustness function we must find maximum failure probability.

¶ The maximum on the upper tail is:

$$\max_{p \in \mathcal{U}(h, \tilde{p})} \int_{\eta x_c}^{\infty} p(u) \, du = \int_{\eta x_c}^{\infty} \tilde{p}(u)(1+h) \, du \quad (38)$$

$$= (1+h) \left[1 - \Phi\left(\frac{\eta x_c}{\sigma}\right) \right] \quad (39)$$

$\Phi(\cdot)$ is the standard normal probability distribution function.

¶ The maximum on the lower tail is the same, so:

$$\max_{p \in \mathcal{U}(h, \tilde{p})} P_f(p) = 2(1+h) \left[1 - \Phi\left(\frac{\eta x_c}{\sigma}\right) \right] \quad (40)$$

¶ We have assumed that h is small enough so that this is no greater than one. This is assured, for some non-negative h , if the nominal density, $\tilde{p}(u)$, entails acceptable probability of failure, which requires that:

$$2 \left[1 - \Phi\left(\frac{\eta x_c}{\sigma}\right) \right] \leq P_c \quad (41)$$

¶ To find \hat{h} from eq.(28) on p.10, equate eq.(40) to P_c , and solve for h :

$$\hat{h}(P_c) = \frac{P_c}{2 \left[1 - \Phi\left(\frac{\eta x_c}{\sigma}\right) \right]} - 1 \quad (42)$$

2.3 Example: Static Poisson Queuing I

¶ Queuing and timing problems:

- Match server rate to client-arrival rate.
 - Inventory problems: keep stock available and fresh.
 - Digital communications synchronization.
- Tracking random events.

¶ The System:

- Server able to handle r clients per day.
- Clients accumulate during the night; no new clients arrive during working hours.
- n = number of clients waiting in morning.
- Clients arrive randomly and independently with constant mean rate, so n is a Poisson random variable:

$$P_n(\lambda) = \frac{e^{-\lambda} \lambda^n}{n!}, \quad n = 0, 1, \dots \quad (43)$$

¶ Uncertainty:

- λ = average number of clients per day. Non-negative
- $\tilde{\lambda}$ = best estimate of λ .
- λ erratically variable, and represented by fractional-error info-gap model:

Approximately:

$$\left| \frac{\lambda - \tilde{\lambda}}{\tilde{\lambda}} \right| \leq h, \quad h \geq 0 \quad (44)$$

More precisely:

$$\mathcal{U}(h, \tilde{\lambda}) = \left\{ \lambda : \max[0, (1 - h)\tilde{\lambda}] \leq \lambda \leq (1 + h)\tilde{\lambda} \right\}, \quad h \geq 0 \quad (45)$$

¶ The Question:

- Manager does not want:
 - Clients who are not handled on the day of arrival: r too small.
 - Unused client-handling capability: r too large.
- What value of r should be adopted?

¶ **Loss function:**

- Probability of Not Serving s_2 or more clients is:

$$\pi_{\text{ns}}(r, \lambda) = \sum_{n=r+s_2}^{\infty} P_n(\lambda) \quad (46)$$

- Probability of Unused Capacity for handling s_1 or more clients is:

$$\pi_{\text{uc}}(r, \lambda) = \sum_{n=0}^{r-s_1} P_n(\lambda) \quad (47)$$

- The loss function is:

$$\pi_{\ell}(r, \lambda) = \pi_{\text{uc}}(r, \lambda) + \pi_{\text{ns}}(r, \lambda) \quad (48)$$

$$= \sum_{n=0}^{r-s_1} P_n(\lambda) + \sum_{n=r+s_2}^{\infty} P_n(\lambda) \quad (49)$$

$$= 1 - \sum_{n=r-s_1+1}^{r+s_2-1} P_n(\lambda) \quad (50)$$

$$= 1 - e^{-\lambda} \sum_{n=r-s_1+1}^{r+s_2-1} \frac{\lambda^n}{n!} \quad (51)$$

- For instance, if $s_1 = s_2 = 1$:

$$\pi_{\ell}(r, \lambda) = 1 - P_r(\lambda) = 1 - \frac{e^{-\lambda} \lambda^r}{r!} \quad (52)$$

¶ **Performance requirement:**

$$\pi_{\ell}(r, \lambda) \leq \varepsilon \quad (53)$$

¶ **Robustness** of handling-capacity r to uncertainty in arrival rate λ :

$$\hat{h}(r, \varepsilon) = \max \left\{ h : \left(\max_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \pi_{\ell}(r, \lambda) \right) \leq \varepsilon \right\} \quad (54)$$

¶ **Inner maximum** in eq.(54):

$$M(h) = \max_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \pi_\ell(r, \lambda) \quad (55)$$

- $M(h)$ increases as h increases because $\mathcal{U}(h, \tilde{\lambda})$ are nested sets:

$$\frac{dM(h)}{dh} \geq 0 \quad (56)$$

- $\hat{h}(r, \varepsilon)$ is greatest h at which:

$$M(h) \leq \varepsilon \quad (57)$$

- Thus $\hat{h}(r, \varepsilon)$ is greatest solution for h of (see fig. 2):

$$M(h) = \varepsilon \quad (58)$$

- In other words, $M(h)$ is the inverse of $\hat{h}(r, \varepsilon)$:

$$M(h) = \varepsilon \quad \text{if and only if} \quad \hat{h}(r, \varepsilon) = h \quad (59)$$

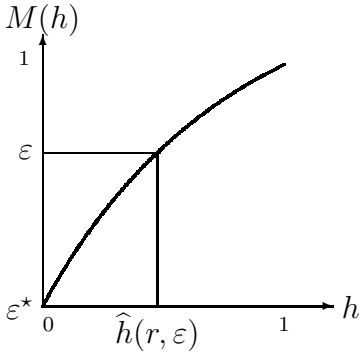


Figure 2: Illustration of the calculation of robustness.

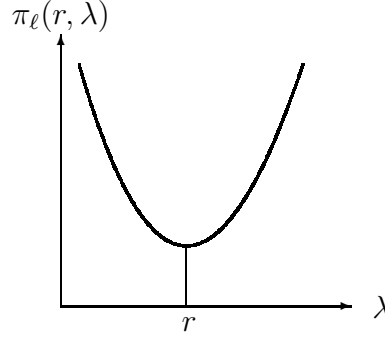


Figure 3: Schematic illustration of $\pi_\ell(r, \lambda)$.

¶ **Evaluating** $M(h)$:

- Consider $s_1 = s_2 = 1$, so $\pi_\ell(r, \lambda)$ in eq.(52), p.14, is:

$$\pi_\ell(r, \lambda) = 1 - \frac{e^{-\lambda} \lambda^r}{r!} \quad (60)$$

- Note, as illustrated schematically in fig. 3, that:

$$\frac{\partial \pi_\ell}{\partial \lambda} = \frac{e^{-\lambda} \lambda^{r-1}}{r!} (\lambda - r) \quad (61)$$

- Hence, $M(h)$ is obtained from eq.(60) with one or the other of the extreme λ values at horizon of uncertainty h . Denote these extreme values:

$$\lambda_+ = (1 + h)\tilde{\lambda} \quad (62)$$

$$\lambda_- = \max[0, (1 - h)\tilde{\lambda}] \quad (63)$$

- Hence:

$$M(h) = \max[\pi_\ell(r, \lambda_-), \pi_\ell(r, \lambda_+)] \quad (64)$$

¶ **Nominal loss function** for $s_1 = s_2 = 1$, eq.(60), p.15:

$$\varepsilon^* = \pi_\ell(r, \tilde{\lambda}) = 1 - \frac{e^{-\tilde{\lambda}} \tilde{\lambda}^r}{r!} \quad (65)$$

This estimate of the loss function is based on the best estimate of the client-arrival rate, $\tilde{\lambda}$.

- Note that:

$$M(0) = \varepsilon^* \quad (66)$$

- Thus, as in fig. 2, p.15:

$$\hat{h}(r, \varepsilon^*) = 0 \quad (67)$$

- The best estimate of the loss function has zero robustness.
- Only worse (larger) loss has positive robustness, as in fig. 2:

$$\varepsilon > \varepsilon' \implies \hat{h}(r, \varepsilon) \geq \hat{h}(r, \varepsilon') \quad (68)$$

¶ **Optimizing the nominal loss function.**

- Optimal server size:

$$r^* = \arg \min_r \pi_\ell(r, \tilde{\lambda}) \quad (69)$$

- Anticipated loss function:

$$\varepsilon^{\text{opt}} = \pi_\ell(r^*, \tilde{\lambda}) \quad (70)$$

- Robustness vanishes as in eq.(67):

$$\hat{h}(r^*, \varepsilon^{\text{opt}}) = 0 \quad (71)$$

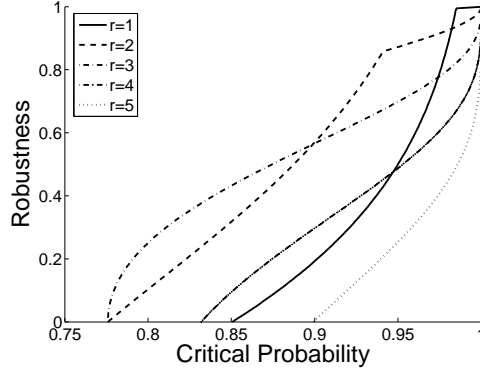


Figure 4: Robustness curves for $\tilde{\lambda} = 3$ and $r = 1, 2, \dots, 5$. $s_1 = s_2 = 1$.

¶ **Numerical example, fig. 4.**

- The best (but highly unreliable) estimate of the number of clients is $\tilde{\lambda} = 3$.
- Fig. 4 shows robustness curves for server-capacities $r = 1, 2, \dots, 5$.
- Recall the loss function, $\pi_\ell(r, \lambda)$, which is the probability of un-served clients or un-used server capacity.
- Consider the loss function at the estimated number of clients, $\pi_\ell(r, \tilde{\lambda})$, which is the x -intersect in fig. 4, shown in table 1:

| r Server capacity | $M(0) = \pi_\ell(r, \tilde{\lambda})$ Nominal loss function |
|---------------------------|---|
| 1 | 0.85 |
| 2 | 0.78 |
| 3 | 0.78 |
| 4 | 0.83 |
| 5 | 0.90 |

Table 1: Nominal loss function for different server capacities.

- We want $\pi_\ell(r, \tilde{\lambda})$ small, so, based on the best-estimate of the client-arrival rate, $\tilde{\lambda}$, our preferences on values of r are:

$$3 \sim_n 2 \succ_n 4 \succ_n 1 \succ_n 5 \tag{72}$$

The subscript ‘n’ indicates that these are ‘nominal’ preferences.

- Now consider the preferences based on the robustness curves, γ_r .
 - An r -value whose curve is further to the right has greater robustness.
 - The following *strict dominances* are observed:

$$3 \gamma_r 4 \gamma_r 5 \tag{73}$$

$$2 \gamma_r 1 \gamma_r 5 \tag{74}$$

- The robust-satisficing preferences in eqs.(73) and (74) are consistent with, but weaker than, the nominal preferences in eq.(72).
- In fig. 4 we see 3 **crossing robustness curves**.
- Crossing of robustness curves implies preference reversal.
- Comparing nominal and robust-satisficing preferences, the differences are shown in table 2:

| γ_n Nominal preference | γ_r robust-satisficing preference |
|-------------------------------------|--|
| $3 \sim_n 2$ | 3 crosses 2 |
| $3 \gamma_n 1$ | 3 crosses 1 |
| $4 \gamma_n 1$ | 4 crosses 1 |

Table 2: Nominal loss function for different server capacities.

- For instance, compare $r = 2$ and $r = 3$ in fig. 4.
 - For $\varepsilon < 0.9$: $\hat{h}(3, \varepsilon) > \hat{h}(2, \varepsilon) \implies 3 \gamma_r 2$.
 - For $\varepsilon > 0.9$: $\hat{h}(2, \varepsilon) > \hat{h}(3, \varepsilon) \implies 2 \gamma_r 3$.
 - Nominally: $3 \sim_n 2$.
- For instance, compare $r = 1$ and $r = 4$ in fig. 4.
 - For $\varepsilon < 0.97$: $\hat{h}(4, \varepsilon) > \hat{h}(1, \varepsilon) \implies 4 \gamma_r 1$.
 - For $\varepsilon > 0.97$: $\hat{h}(1, \varepsilon) > \hat{h}(4, \varepsilon) \implies 1 \gamma_r 4$.
 - Nominally: $4 \sim_n 1$.

2.4 Example: Static Poisson Queuing II

¶ Modify example of section 2.3: different uncertainty in probabilities.

¶ **Uncertain probability distribution:**

- \tilde{P}_n , $n = 0, 1, \dots$ is the best estimated distribution of number of clients accumulated during the night.

- \tilde{P}_n may be Poisson with specified average rate $\tilde{\lambda}$.

- P_n , $n = 0, 1, \dots$ is the unknown actual distribution of number of clients accumulated during the night.

- The info-gap model for P_n is:

$$U(h, \tilde{P}) = \left\{ P_n = \tilde{P}_n + u_n : \max[-\tilde{P}_n, -h\tilde{P}_n] \leq u_n \leq h\tilde{P}_n, \sum_{n=0}^{\infty} u_n = 0 \right\}, \quad h \geq 0 \quad (75)$$

¶ **Loss function:**

- Probability of Not Serving s_2 or more clients is:

$$\pi_{\text{ns}}(r, P) = \sum_{n=r+s_2}^{\infty} (\tilde{P}_n + u_n) \quad (76)$$

- Probability of Unused Capacity for handling s_1 or more clients is:

$$\pi_{\text{uc}}(r, P) = \sum_{n=0}^{r-s_1} (\tilde{P}_n + u_n) \quad (77)$$

- The loss function is:

$$\pi_{\ell}(r, P) = \pi_{\text{uc}}(r, P) + \pi_{\text{ns}}(r, P) \quad (78)$$

$$= \sum_{n=0}^{r-s_1} (\tilde{P}_n + u_n) + \sum_{n=r+s_2}^{\infty} (\tilde{P}_n + u_n) \quad (79)$$

$$= 1 - \sum_{n=r-s_1+1}^{r+s_2-1} (\tilde{P}_n + u_n) \quad (80)$$

- For instance, if $s_1 = s_2 = 1$:

$$\pi_{\ell}(r, P) = 1 - \tilde{P}_r - u_r \quad (81)$$

¶ **Performance requirement**, as before in eq.(53), p.14:

$$\pi_\ell(r, P) \leq \varepsilon \quad (82)$$

¶ **Robustness** of handling-capacity r to uncertainty in arrival rate λ , as in eq.(54), p.14:

$$\hat{h}(r, \varepsilon) = \max \left\{ h : \left(\max_{P \in \mathcal{U}(h, \tilde{P})} \pi_\ell(r, P) \right) \leq \varepsilon \right\} \quad (83)$$

¶ **Inner maximum** in eq.(83):

- Suppose $h \leq 1$ and $\tilde{P}_r \leq 0.5$.
- Then inner maximum occurs for:

$$u_r = -h\tilde{P}_r \quad (84)$$

- Denote inner maximum as $M(h)$, as in eq.(55), p.15.
- Thus, from eq.(81) on p.19:

$$M(h) = 1 - \tilde{P}_r + h\tilde{P}_r = \varepsilon \quad (85)$$

- Robustness is:

$$\hat{h}(r, \varepsilon) = \begin{cases} 0 & \text{if } \varepsilon - 1 + \tilde{P}_r < 0 \\ \frac{\varepsilon - 1 + \tilde{P}_r}{\tilde{P}_r} & \text{else} \end{cases} \quad (86)$$

¶ **Trade-off** of robustness vs. performance, like eq.(68), p.16:

$$\varepsilon > \varepsilon' \implies \hat{h}(r, \varepsilon) \geq \hat{h}(r, \varepsilon') \quad (87)$$

¶ **No robustness of estimated loss**, like eq.(67), p.16:

$$\varepsilon^* = \pi_\ell(r, \tilde{P}) = 1 - \tilde{P}_r \implies \hat{h}(r, \varepsilon^*) = 0 \quad (88)$$

¶ **Robustness function**, eq.(86), p.20, and fig. 5:

- $\hat{h}(r, \varepsilon)$ vs. ε is straight increasing line.

- Two points on the curve are:

$$\hat{h}(r, 1 - \tilde{P}_r) = 0.$$

$$\hat{h}(r, 1) = 1.$$

- Hence:

- Robustness curves cross only at maximal robustness.

- Nominal preference agrees with robust-satisficing preference.

- $\hat{h}(r, \varepsilon)$ quantifies reliability of sub-optimal performance ($\varepsilon > \varepsilon^*$).

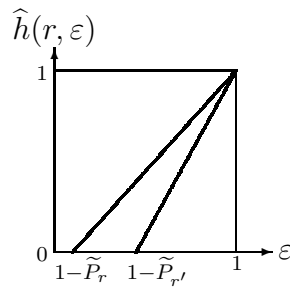


Figure 5: Illustration of robustness curves, eq.(86).

2.5 Example: Dynamic Queuing; Birth and Death Process

¶ Formulation

- Server acts while queue is active.
- n = length of queue of clients waiting for service.
- n can be:
 - positive, meaning that clients are waiting for service.
 - negative, meaning that the server is idle.
 - Thus n can be any integer from $-\infty$ to $+\infty$.
 - Note approximation at both extremes.
- $P_n(t)$ = probability that the length is n at time t .

¶ Birth and death process: differential equations for $P_n(t)$.

- Client arrivals and “departures” are statistically independent.
- λdt = probability of 1 client added during dt .
 λ is uncertain.
- μdt = probability of 1 client removed during dt .
 μ is under our control: client-processing rate of server.
- $1 - \lambda dt - \mu dt$ = probability of 0 clients added or removed during dt .
- Probability-balance equation for $P_n(t)$:

$$P_n(t + dt) = P_n(t)(1 - \lambda dt - \mu dt) + P_{n-1}(t)\lambda dt + P_{n+1}(t)\mu dt + \mathcal{O}(dt^2) + \dots \quad (89)$$

- Re-arrange, divide by dt , take limit $dt \rightarrow 0$:

$$\frac{dP_n(t)}{dt} = \lambda P_{n-1}(t) - \lambda P_n(t) + \mu P_{n+1}(t) - \mu P_n(t), \quad n \in (-\infty, +\infty) \quad (90)$$

- Initial queue size, at $t = 0$, is n_0 , so initial conditions for eqs.(90) are:

$$P_n(0) = \delta_{n_0, n} \quad (91)$$

¶ Moments of $n(t)$:

$$E[n^k(t)] = \sum_{n=-\infty}^{\infty} n^k P_n(t) \quad (92)$$

In particular:

$$\bar{n}(t) = E[n(t)] = \sum_{n=-\infty}^{\infty} n P_n(t) \quad (93)$$

¶ Moment generating function:

- Definition:

$$G(z, t) = \sum_n z^n P_n(t) \quad (94)$$

- Derivative:

$$\frac{\partial G(z, t)}{\partial z} = \sum_n n z^{n-1} P_n(t) \quad (95)$$

- Mean queue size:

$$\left. \frac{\partial G(z, t)}{\partial z} \right|_{z=1} = \sum_n n P_n(t) = E[n(t)] \quad (96)$$

¶ Deriving $G(z, t)$:

- Multiply eq.(90), p.22, by z_n and sum on n over $(-\infty, +\infty)$:

$$\sum_n z^n P'_n = \lambda \sum_n z^n P_{n-1} - (\lambda + \mu) \sum_n z^n P_n + \mu \sum_n z^n P_{n+1} \quad (97)$$

$$= \lambda z \sum_n z^{n-1} P_{n-1} - (\lambda + \mu) \sum_n z^n P_n + \frac{\mu}{z} \sum_n z^{n+1} P_{n+1} \quad (98)$$

$$\frac{\partial G(z, t)}{\partial t} = \lambda z G - (\lambda + \mu) G + \frac{\mu}{z} G \quad (99)$$

$$= \left(\lambda z - (\lambda + \mu) + \frac{\mu}{z} \right) G \quad (100)$$

- Initial condition on $G(z, t)$ at $t = 0$, based on eq.(91), p.22:

$$G(z, t = 0) = z^{n_0} \quad (101)$$

- Integrate eq.(100) on t :

$$G(z, t) = z^{n_0} \exp \left[\left(\lambda z - (\lambda + \mu) + \frac{\mu}{z} \right) t \right] \quad (102)$$

¶ Mean queue size:

Use eqs.(96) and (102) to find:

$$\bar{n}(t, \lambda) = (\lambda - \mu)t + n_0 \quad (103)$$

¶ Uncertainty in λ :

$$\mathcal{U}(h, \tilde{\lambda}) = \left\{ \lambda : \max[0, (1-h)\tilde{\lambda}] \leq \lambda \leq (1+h)\tilde{\lambda} \right\}, \quad h \geq 0 \quad (104)$$

¶ Performance requirement:

$$n_1 \leq \bar{n}(t_c) \leq n_2 \quad (105)$$

- where n_1, n_2 and t_c are specified. Typically, $n_1 < 0$ and $n_2 > 0$.
- t_c is a clearing time chosen by the designer.
- Denote the performance specification $s = (n_1, n_2)$.
- Denote the design variables $q = (\mu, t_c)$.

¶ Robustness with design variables q and specifications s :

$$\hat{h}(q, s) = \max \left\{ h : \left(\max_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \bar{n}(t_c, \lambda) \right) \leq n_2 \text{ and } \left(\min_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \bar{n}(t_c, \lambda) \right) \geq n_1 \right\} \quad (106)$$

¶ Sub-problem robustnesses:

$$\hat{h}_1(q, s) = \max \left\{ h : \left(\min_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \bar{n}(t_c, \lambda) \right) \geq n_1 \right\} \quad (107)$$

$$\hat{h}_2(q, s) = \max \left\{ h : \left(\max_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \bar{n}(t_c, \lambda) \right) \leq n_2 \right\} \quad (108)$$

Since both requirements are necessary:

$$\hat{h}(q, s) = \min[\hat{h}_1(q, s), \hat{h}_2(q, s)] \quad (109)$$

¶ Deriving \hat{h}_2 :

$$\max_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} [(\lambda - \mu)t_c + n_0] \leq n_2 \implies [(1+h)\tilde{\lambda} - \mu]t_c + n_0 \leq n_2 \quad (110)$$

Thus:

$$\hat{h}_2(q, s) = \begin{cases} \frac{n_2 - n_0}{\tilde{\lambda}t_c} + \frac{\mu}{\tilde{\lambda}} - 1 & \text{if } (\tilde{\lambda} - \mu)t_c + n_0 \leq n_2 \\ 0 & \text{else} \end{cases} \quad (111)$$

¶ Deriving \hat{h}_1 :

- The inner minimum in eq.(107) is a decreasing function of h (fig. 6):

$$\min_{\lambda \in \mathcal{U}(h, \tilde{\lambda})} \bar{n}(t_c, \lambda) = \begin{cases} [(1-h)\tilde{\lambda} - \mu] t_c + n_0 & \text{if } h \leq 1 \\ -\mu t_c + n_0 & \text{else} \end{cases} \quad (112)$$

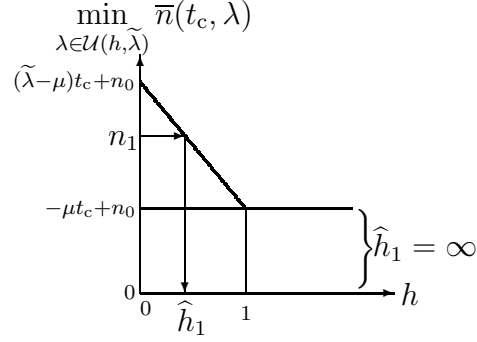


Figure 6: Schematic illustration of the evaluation of \hat{h}_1 from eq.(112).

- Thus:

$$\hat{h}_1(q, s) = \begin{cases} 0 & \text{if } (\tilde{\lambda} - \mu)t_c + n_0 \leq n_1 \\ 1 - \frac{n_1 - n_0}{\tilde{\lambda}t_c} - \frac{\mu}{\tilde{\lambda}} & \text{if } -\mu t_c + n_0 \leq n_1 < (\tilde{\lambda} - \mu)t_c + n_0 \\ \infty & \text{if } n_1 < -\mu t_c + n_0 \end{cases} \quad (113)$$

¶ $\hat{h}(q, s)$ from combining eqs.(109), (111) and (113).

¶ **Maximal robustness.**

• From eq.(109), p.24, we see that the choice of $q = (\mu, t_c)$ which maximizes $\hat{h}(q, s)$ is the choice which causes:

$$\hat{h}_1(q, s) = \hat{h}_2(q, s) \quad (114)$$

- Suppose that n_1 and n_2 are such that $\hat{h}_1(q, s)$ and $\hat{h}_2(q, s)$ are both positive and finite.
- Then eq.(114) is:

$$1 - \frac{n_1 - n_0}{\tilde{\lambda} t_c} - \frac{\mu}{\tilde{\lambda}} = \frac{n_2 - n_0}{\tilde{\lambda} t_c} + \frac{\mu}{\tilde{\lambda}} - 1 \quad (115)$$

which implies:

$$\mu = \tilde{\lambda} + \frac{\Delta}{t_c} \quad \text{where} \quad \Delta = n_0 - \frac{n_1 + n_2}{2} \quad (116)$$

- That is, for any t_c , choosing μ according to eq.(116) maximizes $\hat{h}(q, s)$ for that t_c .
- For any t_c , the robustness with μ from eq.(116) is:

$$\hat{h}(q, s) = \hat{h}_1(q, s) = \hat{h}_2(q, s) = \frac{n_2 - n_1}{2\tilde{\lambda} t_c} \quad (117)$$

provided that n_1 and n_2 are such that $\hat{h}_1(q, s)$ and $\hat{h}_2(q, s)$ are both positive and finite.

- We see from eq.(117) the following trade-offs:

◦ Robustness increases as acceptable un-used capacity increases (as n_1 becomes more negative):

$$\frac{\partial \hat{h}(q, s)}{\partial n_1} < 0 \quad (118)$$

- Robustness increases as the acceptable # of un-served clients increases:

$$\frac{\partial \hat{h}(q, s)}{\partial n_2} > 0 \quad (119)$$

- Robustness increases as the tolerance-window $n_2 - n_1$ increases:

$$\frac{\partial \hat{h}(q, s)}{\partial (n_2 - n_1)} > 0 \quad (120)$$

- Robustness increases as clearing time decreases:

$$\frac{\partial \hat{h}(q, s)}{\partial t_c} < 0 \quad (121)$$

3 Probabilistic Info-Gap Parameter

¶ Basic idea:

- Complex temporal or spatial waveforms are modelled by an info-gap model, $\mathcal{U}(h, \tilde{u})$, $h \geq 0$.
- The uncertainty parameter h has physical meaning. E.g. energy of event.
- The uncertainty in h is represented by a pdf.

¶ Example:

- Dynamic system with uncertain load $u \in \mathcal{U}(h, \tilde{u})$, $h \geq 0$.
- Load u causes damage $\delta(u)$.
- Failure if:

$$\delta_u(t) \geq \Delta_c \quad (122)$$

¶ Robustness:

$$\hat{h}(q, \Delta_c) = \max \left\{ h : \left(\max_{u \in \mathcal{U}(h, \tilde{u})} \delta_u(t) \right) \leq \Delta_c \right\} \quad (123)$$

q is the vector of decision variables.

¶ Failure **can not occur** if:

$$h < \hat{h}(q, \Delta_c) \quad (124)$$

¶ Failure **need not occur** even if:

$$h \geq \hat{h}(q, \Delta_c) \quad (125)$$

(Load may be propitious.)

¶ We **cannot calculate** P_f because $p(u)$ is unknown.

¶ We **can calculate** an upper bound for P_f :

$$P_f \leq \text{Prob} [h \geq \hat{h}(q, \Delta_c)] = 1 - P [\hat{h}(q, \Delta_c)] \quad (126)$$

$P(\cdot)$ is the cumulative probability distribution of h .

¶ Optimal q :

- We can seek q to maximize $\hat{h}(q, \Delta_c)$.
- $P(h)$ is a monotonically increasing function.
- Thus maximizing $\hat{h}(q, \Delta_c)$ also maximizes $P(\hat{h})$ and minimizes $1 - P(\hat{h})$.

¶ Proof:

$$\partial P(h)/\partial h \geq 0 \quad (127)$$

and because:

$$\frac{\partial P[\hat{h}(q, \Delta_c)]}{\partial q} = \frac{\partial P[\hat{h}(q, \Delta_c)]}{\partial h} \frac{\partial \hat{h}(q, \Delta_c)}{\partial q} \quad (128)$$

QED

¶ Equivalent definition of the robust optimal action \hat{q} :

$$\hat{h}(\hat{q}, \Delta_c) = \max_{q \in \mathcal{Q}} P[\hat{h}(q, \Delta_c)] \quad (129)$$

¶ Likewise, $P(\cdot)$ defines the same preference ordering on q as $\hat{h}(q, \Delta_c)$:

$$q \succ q' \quad \text{if} \quad P[\hat{h}(q, \Delta_c)] > P[\hat{h}(q', \Delta_c)] \quad (130)$$

¶ This provides a probabilistic calibration of the relative merits of the options.